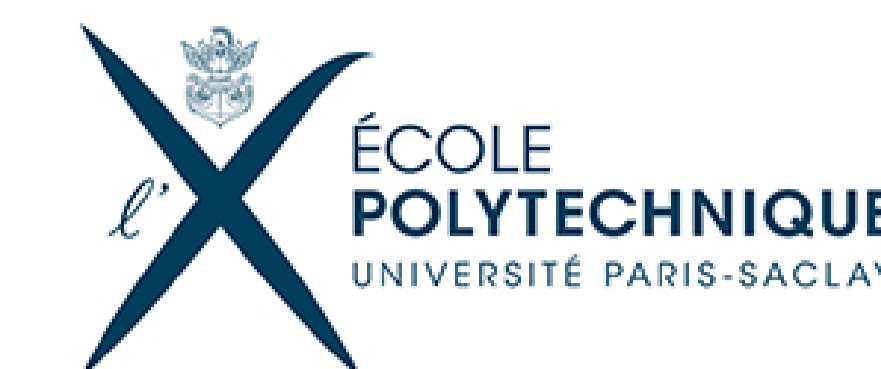


# Non-asymptotic Analysis of Biased Stochastic Approximation Scheme

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## Stochastic Approximation

- **Objective:** Find a *stationary point* of smooth Lyapunov function  $V(\eta)$ .
- SA scheme (Robbins and Monro, 1951) is a stochastic process:

$$\eta_{n+1} = \eta_n - \gamma_{n+1} H_{\eta_n}(X_{n+1}), \quad n \in \mathbb{N} \quad (1)$$

where  $\eta_n \in \mathcal{H} \subseteq \mathbb{R}^d$  is the  $n$ th state,  $\gamma_n > 0$  is the step size.

- The *drift term*  $H_{\eta_n}(X_{n+1})$  depends on an **i.i.d. random element**  $X_{n+1}$  and

$$h(\eta_n) = \mathbb{E}[H_{\eta_n}(X_{n+1}) | \mathcal{F}_n] = \nabla V(\eta_n),$$

where  $\mathcal{F}_n = \sigma(\eta_0, \{X_m\}_{m \leq n})$ . In this case, SA is better known as the SGD method.

## Biased SA Scheme

- The **mean field is biased**  $\Leftarrow$  gradient is sometimes difficult to compute...

We have  $h(\eta) \neq \nabla V(\eta)$  and for some  $c_0 \geq 0, c_1 > 0$ ,

$$c_0 + c_1 \langle \nabla V(\eta) | h(\eta) \rangle \geq \|h(\eta)\|^2, \quad \forall \eta \in \mathcal{H}$$

- The **drift term**  $\{H_{\eta_n}(X_{n+1})\}_{n \geq 1}$  is **not i.i.d.**. For example, in reinforcement learning,  $\eta_n$  controls the policy in a MDP &  $H_{\eta_n}(X_{n+1})$  is computed from the MDP's state.

The random elements  $\{X_n\}_{n \geq 1}$  form a **state-dependent Markov chain**:

$$\mathbb{E}[H_{\eta_n}(X_{n+1}) | \mathcal{F}_n] = P_{\eta_n} H_{\eta_n}(X_n) = \int H_{\eta_n}(x) P_{\eta_n}(X_n, dx),$$

where  $P_{\eta_n} : X \times \mathcal{X} \rightarrow \mathbb{R}_+$  is Markov kernel with a unique stationary distribution  $\pi_{\eta_n}$ .

- In the latter case, the mean field is given by  $h(\eta) = \int H_{\eta}(x) \pi_{\eta}(dx)$ .

- **Stopping criterion:** fix any  $n \geq 1$ , we stop the SA at a random iteration  $N$  with

$$\mathbb{P}(N = \ell) = \left( \sum_{k=0}^{\ell-1} \gamma_{k+1} \right)^{-1} \gamma_{\ell+1}, \quad \text{with } N \in \{1, \dots, n\}.$$

## Prior Work

- We focus on the **non-asymptotic convergence** analysis of SA scheme, where the relevant results are rare. Define:

$$e_{n+1} := H_{\eta_n}(X_{n+1}) - h(\eta_n) \quad (2)$$

**Case 1: When  $\{e_n\}_{n \geq 1}$  is Martingale difference**  $\rightarrow \mathbb{E}[e_{n+1} | \mathcal{F}_n] = 0$

- *Asymptotic analysis:* (Robbins and Monro, 1951); *Non-asymptotic analysis:* (Ghadimi and Lan, 2013).

**Case 2: When  $\{e_n\}_{n \geq 1}$  is state-controlled Markov noise**

$$\mathbb{E}[e_{n+1} | \mathcal{F}_n] = P_{\eta_n} H_{\eta_n}(X_n) - h(\eta_n) \neq 0.$$

- *Asymptotic analysis:* (Tadić and Doucet, 2017); *Non-asymptotic analysis:* (Sun et al., 2018), (Duchi et al., 2012), (Bhandari et al., 2018)

## Analysis For Martingale Difference Noise (Case 1)

**Assumption:**  $\mathbb{E}[e_{n+1} | \mathcal{F}_n] = 0$ ,  $\mathbb{E}[\|e_{n+1}\|^2 | \mathcal{F}_n] \leq \sigma_0^2 + \sigma_1^2 \|h(\eta_n)\|^2$ . (e.g., when  $X_n$  is i.i.d. similar to the SGD setting).

**Theorem 1.** Let  $\gamma_{n+1} \leq (2c_1 L(1 + \sigma_1^2))^{-1}$  and  $V_{0,n} := \mathbb{E}[V(\eta_0) - V(\eta_{n+1})]$ ,

$$\mathbb{E}[\|h(\eta_N)\|^2] \leq \frac{2c_1(V_{0,n} + \sigma_0^2 L \sum_{k=0}^n \gamma_{k+1}^2)}{\sum_{k=0}^n \gamma_{k+1}} + 2c_0,$$

Set  $\gamma_k = (2c_1 L(1 + \sigma_1^2) \sqrt{k})^{-1} \Rightarrow \mathbb{E}[\|h(\eta_N)\|^2] = \mathcal{O}(c_0 + \log n / \sqrt{n})$ . Remark: if  $h(\eta) = \nabla V(\eta)$  (with  $c_0 = d_0 = 0$ ), it recovers (Ghadimi and Lan, 2013, Theorem 2.1).

## Analysis For State-dependent Markov Noise (Case 2)

**Assumptions:** we need a few regularity conditions in this case,

1. There exists a Borel measurable function  $\hat{H} : \mathcal{H} \times X \rightarrow \mathcal{H}$ ,

$$\hat{H}_{\eta}(x) - P_{\eta} \hat{H}_{\eta}(x) = H_{\eta}(x) - h(\eta), \quad \forall \eta \in \mathcal{H}, x \in X.$$

$\Rightarrow$  existence of solution to the **Poisson equation**.

2. For all  $\eta \in \mathcal{H}$  and  $x \in X$ ,  $\|\hat{H}_{\eta}(x)\| \leq L_{PH}^{(0)}$ ,  $\|P_{\eta} \hat{H}_{\eta}(x)\| \leq L_{PH}^{(0)}$ , and

$$\sup_{x \in X} \|P_{\eta} \hat{H}_{\eta}(x) - P_{\eta'} \hat{H}_{\eta'}(x)\| \leq L_{PH}^{(1)} \|\eta - \eta'\|, \quad \forall (\eta, \eta') \in \mathcal{H}^2.$$

$\Rightarrow$  **smoothness** of  $\hat{H}_{\eta}(x)$ , satisfied if  $P_{\eta}, H_{\eta}(X)$  are smooth w.r.t.  $\eta$ .

3. It holds that  $\sup_{\eta \in \mathcal{H}, x \in X} \|H_{\eta}(x) - h(\eta)\| \leq \sigma$ .

$\Rightarrow$  requires the noise is **uniformly bounded** for all  $x \in X$ .

**Example:** assumptions 1 & 2 are satisfied if the Markov kernel  $P_{\eta_n}$  is geometrically ergodic + smooth, and the drift term is smooth w.r.t.  $\eta$ .

**Theorem 2.** Suppose that the step sizes are decreasing and  $\gamma_1 \leq 0.5(c_1(L + C_h))^{-1}$  (+other conditions). Let  $V_{0,n} := \mathbb{E}[V(\eta_0) - V(\eta_{n+1})]$ ,

$$\mathbb{E}[\|h(\eta_N)\|^2] \leq \frac{2c_1(V_{0,n} + C_{0,n} + (\sigma^2 L + C_{\gamma}) \sum_{k=0}^n \gamma_{k+1}^2)}{\sum_{k=0}^n \gamma_{k+1}} + 2c_0.$$

- Set  $\gamma_k = (2c_1 L(1 + C_h) \sqrt{k})^{-1} \Rightarrow \mathbb{E}[\|h(\eta_N)\|^2] = \mathcal{O}(c_0 + \log n / \sqrt{n})$  (same as Case 1).

- **Proof idea:** challenge is that  $e_{n+1}$  is not zero-mean  $\Rightarrow$  bound the sum of  $\mathbb{E}[\langle \nabla V(\eta_n) | e_{n+1} \rangle]$  w/ Poisson equation + a novel decomposition (cf. Lemma 2).

## Regularized Online EM Algorithm

- **Special Case of GMM:** we fit the data  $\{Y_n\}_{n \geq 1}$ ,  $Y_n \sim \pi$  into the parametric model with  $\theta = (\{\omega_m\}_{m=1}^{M-1}, \{\mu_m\}_{m=1}^M)$

$$g(y; \theta) \propto \left(1 - \sum_{m=1}^{M-1} \omega_m\right) \exp\left(-\frac{(y - \mu_M)^2}{2}\right) + \sum_{m=1}^{M-1} \omega_m \exp\left(-\frac{(y - \mu_m)^2}{2}\right),$$

- Data arrives in a streaming fashion, Cappé and Moulines (2009) does:

$$\text{E-step: } \hat{s}_{n+1} = \hat{s}_n + \gamma_{n+1} \{\bar{s}(Y_{n+1}; \hat{\theta}_n) - \hat{s}_n\},$$

$$\text{M-step: } \hat{\theta}_{n+1} = \bar{\theta}(\hat{s}_{n+1}).$$

- The **E-step** is a biased SA step on  $s$  with the drift term & mean field

$$H_{\hat{s}_n}(Y_{n+1}) = \hat{s}_n - \bar{s}(Y_{n+1}; \bar{\theta}(\hat{s}_n)), \quad h(\hat{s}_n) = \hat{s}_n - \mathbb{E}_{\pi}[\bar{s}(Y_{n+1}; \bar{\theta}(\hat{s}_n))]$$

## Analysis of the ro-EM Algorithm (Application of Case 1)

Consider the KL divergence as a function of sufficient statistics  $s$ :

$$V(s) := \text{KL}(\pi | g(\cdot; \bar{\theta}(s))) + R(\bar{\theta}(s)) = \mathbb{E}_{\pi}[\log(\pi(Y)/g(Y; \bar{\theta}(s)))] + R(\bar{\theta}(s)).$$

**Corollary 1.** Set  $\gamma_k = (2c_1 L(1 + \sigma_1^2) \sqrt{k})^{-1}$ . Ro-EM method for GMM finds  $\hat{s}_N$  such that

$$\mathbb{E}[\|\nabla V(\hat{s}_N)\|^2] = \mathcal{O}(\log n / \sqrt{n})$$

The expectation is taken w.r.t.  $N$  and the observation law  $\pi$ .

- First **explicit non-asymptotic** rate given for online EM method.

- Consider a slightly modified/regularized M-step update for satisfaction of the technical conditions.

## (Online) Policy Gradient Method

- Consider a Markov Decision Process (MDP)  $(S, A, R, P)$ :  
–  $S, A$  is the finite set of state/action.  
–  $R : S \times A \rightarrow [0, R_{\max}]$  is a reward function;  $P$  is the transition model.
- A **policy** is parameterized by  $\eta \in \mathbb{R}^d$  as (e.g., soft-max):

$$\Pi_{\eta}(a'; s') = \text{probability of taking action } a' \text{ in state } s'$$

- Update  $\eta$  in an online fashion (Tadić and Doucet, 2017) using observed state-action pair:

$$G_{n+1} = \lambda G_n + \nabla \log \Pi_{\eta_n}(A_{n+1}; S_{n+1}),$$

$$\eta_{n+1} = \eta_n + \gamma_{n+1} G_{n+1} R(S_{n+1}, A_{n+1})$$

where  $\lambda \in (0, 1)$  is a parameter for the variance-bias trade-off.

- The  $\eta$ -update is an biased SA step with the drift term:

$$H_{\eta_n}(X_{n+1}) = G_{n+1} R(S_{n+1}, A_{n+1})$$

## Analysis of Policy Gradient Method (Application of Case 2)

Let  $\nu_{\eta}(s, a)$  be the invariant distribution of  $\{(S_t, A_t)\}_{t \geq 1}$ , we consider:

$$J(\eta) := \sum_{s \in S, a \in A} \nu_{\eta}(s, a) R(s, a).$$

**Corollary 2.** Set  $\gamma_k = (2c_1 L(1 + C_h) \sqrt{k})^{-1}$ . For any  $n \in \mathbb{N}$ , the policy gradient algorithm (3) finds a policy that

$$\mathbb{E}[\|\nabla J(\eta_N)\|^2] = \mathcal{O}\left((1 - \lambda)^2 \Gamma^2 + c(\lambda) \log n / \sqrt{n}\right),$$

where  $c(\lambda) = \mathcal{O}(\frac{1}{1-\lambda})$ . Expectation is taken w.r.t.  $N$  and  $(A_n, S_n)$ .

- It shows the **first convergence rate** for the online PG method.
- Our result shows the **variance-bias trade-off** with  $\lambda \in (0, 1)$ .
- Setting  $\lambda \rightarrow 1$  reduces the bias, but decreases the convergence speed.

## Conclusion

- **Theorem 1 & 2** show the non-asymptotic convergence rate of biased SA scheme with smooth (possibly non-convex) Lyapunov function.
- With appropriate step size, in  $n$  iterations the SA scheme finds  $\mathbb{E}[\|h(\eta_N)\|^2] = \mathcal{O}(c_0 + \log n / \sqrt{n})$ , where  $c_0$  is the bias and  $h(\cdot)$  is the mean field.
- Applications to online EM and online policy gradient.

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