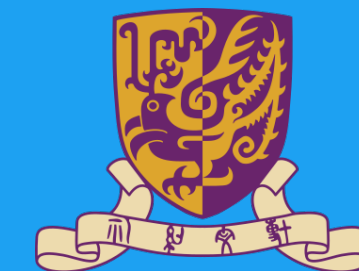


# On the Global Convergence of (Fast) Incremental EM Methods

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# Maximum Likelihood Estimation (MLE)

- We have vectors of data  $Y$  that are *observed* and  $Z$  that are *latent*
- We assume a probabilistic model on the observations  $Y$ ,  $g(Y, \theta)$
- We can define  $f(Z, Y, \theta)$  as the complete data likelihood and  $p(Z|Y, \theta)$  as the conditional distribution of  $Z$  given  $Y$
- The MLE problem is, given a model  $g(Y, \theta)$  and some actual data  $Y$ , find the parameter  $\theta$  which makes the data most likely:

$$\theta^{ML} := \arg \max_{\theta} g(Y, \theta)$$

- This problem is an **optimization problem**, which we could use any imaginable tool to solve
- In practice, it's often **hard** to get expressions for the **derivatives** needed by **gradient** methods
- **Expectation-Maximization (EM)** method is one popular and powerful way of proceeding, but not the only way. **It takes advantage of the latent data to complete the observations.**

# Context

## Settings and Notations

- Many problems in machine learning pertain to tackling an empirical risk minimization of the form

$$\min_{\theta \in \Theta} \bar{\mathcal{L}}(\theta) := \mathcal{L}(\theta) + R(\theta) \quad \text{with} \quad \mathcal{L}(\theta) = \frac{1}{n} \sum_{i=1}^n \mathcal{L}_i(\theta) := \frac{1}{n} \sum_{i=1}^n \{-\log g(y_i; \theta)\}$$

- $\{y_i\}_{i=1}^n$  are the observations,  $\Theta$  is a convex subset of  $\mathbb{R}^d$ ,  $R(\theta)$  is a smooth convex regularization function.
- The objective function  $\bar{\mathcal{L}}(\theta)$  is possibly **nonconvex** and is assumed to be **lower bounded**  $\bar{\mathcal{L}}(\theta) > -\infty$

## Exponential Family

- Latent data model:  $\{z_i\}_{i=1}^n$  are not observed
- Complete data likelihood belongs to the curved exponential family:

Sufficient statistics takes values in  $S \subset \mathbb{R}^d$

$$f(z_i, y_i; \theta) = h(z_i, y_i) \exp(\langle S(z_i, y_i) | \phi(\theta) \rangle - \psi(\theta))$$

# EM Method for Exponential Family

## Updates

### ▸ E-step:

$$\bar{s}(\theta) = \frac{1}{n} \sum_{i=1}^n \bar{s}_i(\theta)$$

where:

$$\bar{s}_i(\theta) = \int_{\mathbf{Z}} S(z_i, y_i) p(z_i | y_i; \theta) \mu(dz_i)$$

### ▸ Define the function $L(\cdot; \theta) : S \rightarrow \mathbb{R}$ as:

$$L(s; \theta) := R(\theta) + \psi(\theta) - \langle s | \phi(\theta) \rangle$$

### ▸ There exists a function $\bar{\theta} : S \mapsto \Theta$ such that

$$L(s; \bar{\theta}(s)) \leq L(s; \theta)$$

### ▸ M-step:

$$\theta = \bar{\theta}(\bar{s}) = \arg \min_{\theta \in \Theta} \{R(\theta) + \psi(\theta) - \langle s | \phi(\theta) \rangle\}$$

## Limitations

- Even though the EM has appealing features:
  - Monotone in likelihood
  - Invariant w.r.t. parametrization
  - Numerically stable (well defined set)
- It is not applicable with the sheer size of today's data
- Approaches based on Stochastic Optimization:
  - [Neal and Hinton, 1998]: Incremental EM (iEM)
  - [Cappé and Moulines, 2009]: Online EM (sEM)
  - [Chen+, 2018]: Variance Reduces EM (sEM-VR)



# Stochastic Optimization for EM Methods

## General Formulation

▸ Stochastic EM:

**sE-step:**  $\hat{\mathbf{s}}^{(k+1)} = \hat{\mathbf{s}}^{(k)} - \gamma_{k+1} \left( \hat{\mathbf{s}}^{(k)} - \mathcal{S}^{(k+1)} \right)$

where  $\gamma_k$  is the stepsize and  $\mathcal{S}^{(k+1)}$  is a proxy for  $\bar{\mathbf{s}} \left( \boldsymbol{\theta}^{(k)} \right)$

▸ **M-step:**

$$\boldsymbol{\theta}^{(k+1)} = \bar{\boldsymbol{\theta}}(\hat{\mathbf{s}}^{(k+1)}) = \arg \min_{\boldsymbol{\theta} \in \Theta} \{ R(\boldsymbol{\theta}) + \psi(\boldsymbol{\theta}) - \langle \hat{\mathbf{s}}^{(k+1)} | \phi(\boldsymbol{\theta}) \rangle \}$$

▸ We simplify the notations:

$$\bar{\mathbf{s}}_i^{(k)} := \bar{\mathbf{s}}_i \left( \boldsymbol{\theta}^{(k)} \right) = \int_{\mathbf{Z}} S(z_i, y_i) p \left( z_i | y_i; \hat{\boldsymbol{\theta}}^{(k)} \right) \mu(\mathrm{d}z_i)$$

$$\bar{\mathbf{s}}^{(k)} := \bar{\mathbf{s}} \left( \boldsymbol{\theta}^{(k)} \right) = \frac{1}{n} \sum_{i=1}^n \bar{\mathbf{s}}_i^{(k)}$$

$$\ell(k) := m \lfloor k/m \rfloor \quad \text{First iteration number of the current epoch}$$

$$(iEM [NH, 1998]) \quad \mathcal{S}^{(k+1)} = \mathcal{S}^{(k)} + \frac{1}{n} (\bar{\mathbf{s}}_{i_k}^{(k)} - \bar{\mathbf{s}}_{i_k}^{(\tau_{i_k}^k)}) \quad [1]$$

$$(sEM [CM, 2009]) \quad \mathcal{S}^{(k+1)} = \bar{\mathbf{s}}_{i_k}^{(k)} \quad [2]$$

$$(sEM - VR [CZTZ., 2018]) \quad \mathcal{S}^{(k+1)} = \bar{\mathbf{s}}^{(\ell(k))} + (\bar{\mathbf{s}}_{i_k}^{(k)} - \bar{\mathbf{s}}_{i_k}^{(\ell(k))}) \quad [3]$$

$$(fiEM [KLMW., 2019]) \quad \begin{aligned} \mathcal{S}^{(k+1)} &= \bar{\mathcal{S}}^{(k)} + (\bar{\mathbf{s}}_{i_k}^{(k)} - \bar{\mathbf{s}}_{i_k}^{(t_{i_k}^k)}) \\ \bar{\mathcal{S}}^{(k+1)} &= \bar{\mathcal{S}}^{(k)} + n^{-1} (\bar{\mathbf{s}}_{j_k}^{(k)} - \bar{\mathbf{s}}_{j_k}^{(t_{j_k}^k)}). \end{aligned} \quad [4]$$

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### Algorithm 3 sEM algorithms

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**Initialization:** initializations  $\hat{\boldsymbol{\theta}}^{(0)} \leftarrow 0$ ,  $\hat{\mathbf{s}}^{(0)} \leftarrow \bar{\mathbf{s}}^{(0)}$ ,  $K_{\max} \leftarrow \text{max. iteration number}$ .

Set the terminating iteration number,  $K \in \{0, \dots, K_{\max} - 1\}$ , as a discrete r.v. with:

$$P(K = k) = \frac{\gamma_k}{\sum_{\ell=0}^{K_{\max}-1} \gamma_{\ell}}. \quad (42)$$

**Iteration k:** Given the current state of the chain  $\psi_i^{(t-1)}$ :

1. Draw index  $i_k \in \llbracket 1, n \rrbracket$  uniformly (and  $j_k \in \llbracket 1, n \rrbracket$  for fiEM).
2. Compute the surrogate sufficient statistics  $\mathcal{S}^{(k+1)}$  using [1] or [2] or [3] or [4]
3. Compute  $\hat{\mathbf{s}}^{(k+1)}$  via the sE-step
4. Compute  $\boldsymbol{\theta}^{(k+1)}$  via the M-step

**Return:**  $\boldsymbol{\theta}^{(K)}$ .

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# Global Convergence

## Assumptions

**(A1)** The function  $\phi$  is smooth and bounded on the interior of  $\Theta$ , noted  $\text{int}(\Theta)$

For all  $(\theta, \theta') \in \text{int}(\Theta)$ ,  $\|J_\phi^\theta(\theta) - J_\phi^\theta(\theta')\| \leq L_\phi \|\theta - \theta'\|$   
and  $\|J_\phi^\theta(\theta')\| \leq C_\phi$

**(A2)** The conditional distribution is smooth on  $\text{int}(\Theta)$

$$|p(z|y_i; \theta) - p(z|y_i; \theta')| \leq L_p \|\theta - \theta'\|$$

**(A3)** The function  $\theta \rightarrow L(s; \theta) := R(\theta) + \psi(\theta) - \langle s | \phi(\theta) \rangle$  admits a unique global minimum

Also,  $J_\phi^\theta(\bar{\theta}(s))$  is full rank and  $\bar{\theta}(s)$  is  $L_\theta$ -Lipschitz

Define:

$$B(s) := J_\phi^\theta(\bar{\theta}(s)) (H_L^\theta(s, \bar{\theta}(s)))^{-1} J_\phi^\theta(\bar{\theta}(s))^\top$$

**(A4)**  $v_{\max} := \sup_{s \in S} \|B(s)\| < \infty$  and  $0 < v_{\min} := \inf_{s \in S} \lambda_{\min}(B(s))$

$$\|B(s) - B(s')\| \leq L_B \|s - s'\|$$

## Incremental EM Method

### Lemma

Under **(A1)-(A4)**, define  $e_i(\theta; \theta') := Q_i(\theta; \theta') - \mathcal{L}_i(\theta)$

We have

$$\|\nabla e_i(\theta; \theta') - \nabla e_i(\bar{\theta}; \theta')\| \leq L_e \|\theta - \bar{\theta}\|$$

where  $L_e := C_\phi C_Z L_p + C_S L_\phi$

### Theorem

Under **(A1)-(A4)** for the iEM **[1]** for any  $K_{\max} \geq 1$

$$\mathbb{E} \left[ \left\| \nabla \bar{\mathcal{L}}(\theta^{(K)}) \right\|^2 \right] \leq n \frac{2L_e}{K_{\max}} \mathbb{E} \left[ \bar{\mathcal{L}}(\theta^{(0)}) - \bar{\mathcal{L}}(\theta^{(K_{\max})}) \right]$$

where  $L_e$  is defined above and  $K$  is a uniform random variable on  $[0, K_{\max} - 1]$  and independent of the  $\{i_k\}_{k=0}^{K_{\max}}$

# Stochastic EM as Scaled Gradient Methods

- From a (Scaled) Gradients Method point of view, we consider the minimization problem:

$$\min_{\mathbf{s} \in \mathcal{S}} V(\mathbf{s}) := \bar{\mathcal{L}}(\bar{\boldsymbol{\theta}}(\mathbf{s})) = R(\bar{\boldsymbol{\theta}}(\mathbf{s})) + \frac{1}{n} \sum_{i=1}^n \mathcal{L}_i(\bar{\boldsymbol{\theta}}(\mathbf{s}))$$

## Lemma

Under **(A1)-(A4)**, we have

$$\|\bar{\mathbf{s}}_i(\bar{\boldsymbol{\theta}}(\mathbf{s})) - \bar{\mathbf{s}}_i(\bar{\boldsymbol{\theta}}(\mathbf{s}'))\| \leq \mathbf{L}_s \|\mathbf{s} - \mathbf{s}'\|$$

$$\|\nabla V(\mathbf{s}) - \nabla V(\mathbf{s}')\| \leq \mathbf{L}_V \|\mathbf{s} - \mathbf{s}'\|$$

where  $\mathbf{L}_s := C_Z \mathbf{L}_p \mathbf{L}_\theta$  and  $\mathbf{L}_V := v_{\max}(1 + \mathbf{L}_s) + \mathbf{L}_B C_S$

## Theorem (sEM-VR)

There exists a constant  $\mu \in (0, 1)$  such that if

$$\bar{L}_v := \max(L_V, L_s) \quad \gamma = \frac{\mu v_{\min}}{\bar{L}_v n^{2/3}} \quad m = \frac{n}{2\mu^2 v_{\min}^2 + \mu}$$

Then:

$$\mathbb{E} \left[ \left\| \nabla V \left( \hat{\mathbf{s}}^{(K)} \right) \right\|^2 \right] \leq n^{\frac{2}{3}} \frac{2\bar{L}_v}{\mu K_{\max}} \frac{v_{\max}^2}{v_{\min}^2} \mathbb{E} \left[ V \left( \hat{\mathbf{s}}^{(0)} \right) - V \left( \hat{\mathbf{s}}^{(K_{\max})} \right) \right]$$

## Theorem (fiEM)

There exists a constant  $\mu \in (0, 1)$  such that if

$$\bar{L}_v := \max(L_V, L_s) \quad \gamma = \frac{v_{\min}}{\alpha \bar{L}_v n^{2/3}} \quad \alpha := \max(6, 1 + 4v_{\min})$$

Then:

$$\mathbb{E} \left[ \left\| \nabla V \left( \hat{\mathbf{s}}^{(K)} \right) \right\|^2 \right] \leq n^{\frac{2}{3}} \frac{\alpha^2 \bar{L}_v}{K_{\max}} \frac{v_{\max}^2}{v_{\min}^2} \mathbb{E} \left[ V \left( \hat{\mathbf{s}}^{(0)} \right) - V \left( \hat{\mathbf{s}}^{(K_{\max})} \right) \right]$$



# Numerical Applications

## Gaussian Mixture Models (GMM)

- Fit a GMM model to a set of  $n$  observations
- Each of  $M$  components with unit variance
- The complete log likelihood reads:

$$\begin{aligned} \log f(z_i, y_i; \boldsymbol{\theta}) &= \sum_{m=1}^M 1_{\{m\}}(z_i) [\log(\omega_m) - \mu_m^2/2] \\ &+ \sum_{m=1}^M 1_{\{m\}}(z_i) \mu_m y_i + \text{constant} \end{aligned}$$

$$\boldsymbol{\theta} := (\boldsymbol{\omega}, \boldsymbol{\mu}) \quad \boldsymbol{\omega} = \{\omega_m\}_{m=1}^{M-1} \quad \boldsymbol{\mu} = \{\mu_m\}_{m=1}^M$$

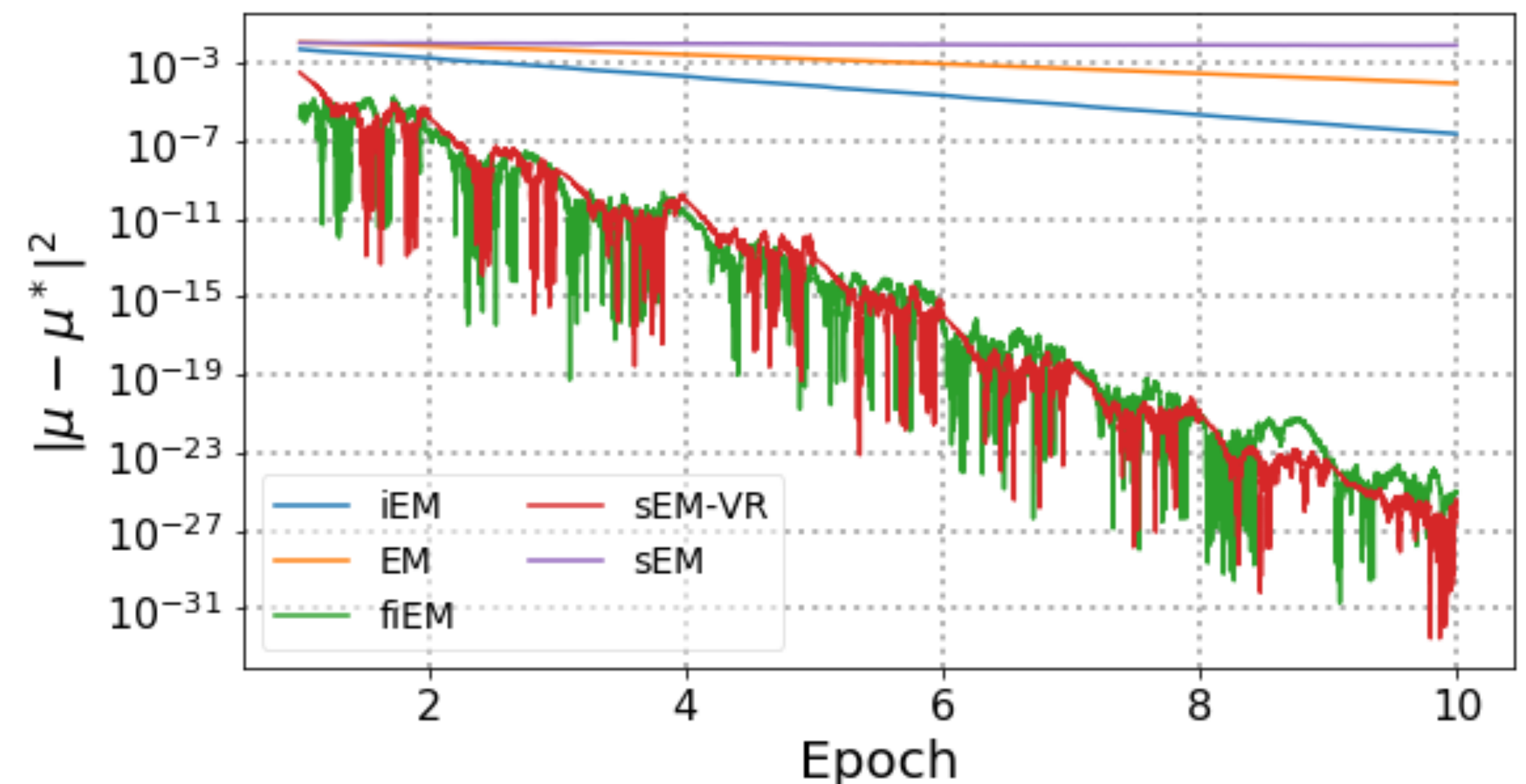
- Penalization used:

$$R(\boldsymbol{\theta}) = \frac{\delta}{2} \sum_{m=1}^M \mu_m^2 - \log \text{Dir}(\boldsymbol{\omega}; M, \epsilon)$$

- Numerical: GMM with  $M=2$  and  $\mu_1 = -\mu_2 = 0.5$

## Experiments

- **Fixed sample size:** size  $n = 10^4$  and run to get  $\mu^*$
- Stepsize for sEM  $\gamma_k = 3/(k + 10)$
- Stepsize for sEM-VR and fiEM prop. to  $1/n^{2/3}$



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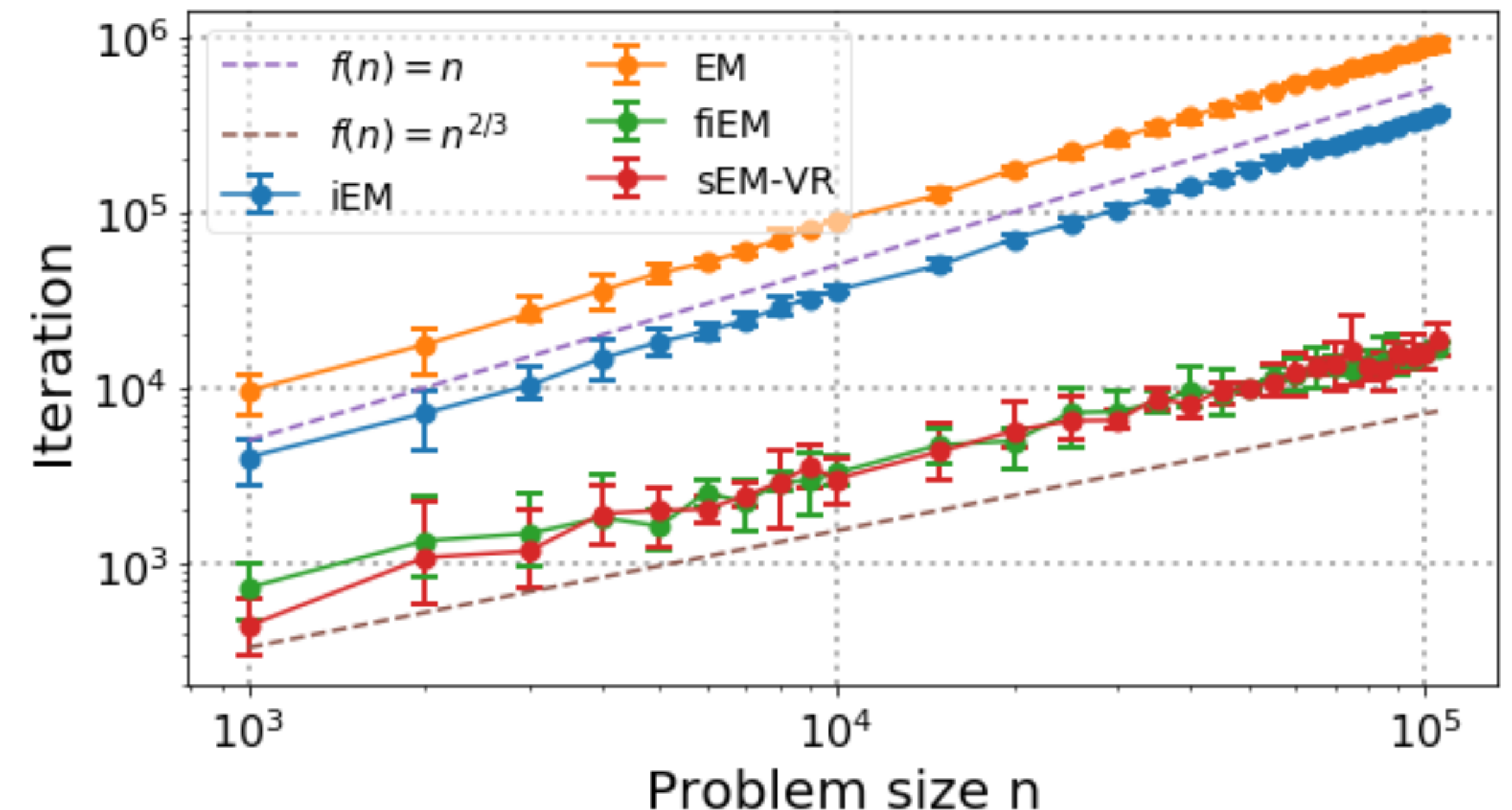
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## Experiments

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Stepsize for sEM  $\gamma_k = 3/(k + 10)$   
Stepsize for sEM-VR and fiEM prop. to  $1/n^{2/3}$
- **Varying sample size:** nb. iterations required to reach a precision of  $10^{-3}$  from  $n = 10^3$  to  $n = 10^5$



# Numerical Applications

## Probabilistic Latent Semantic Analysis

- Consider  $D$  documents with terms from a vocabulary of size  $V$ .
- Data is summarized by a list of tokens

$$\{y_i\}_{i=1}^n \quad y_i = \left( y_i^{(d)}, y_i^{(w)} \right)$$

- The goal of pLSA is to classify the documents into  $K$  topics which is modeled as a latent variable associated with each token  $z_i \in [1, K]$

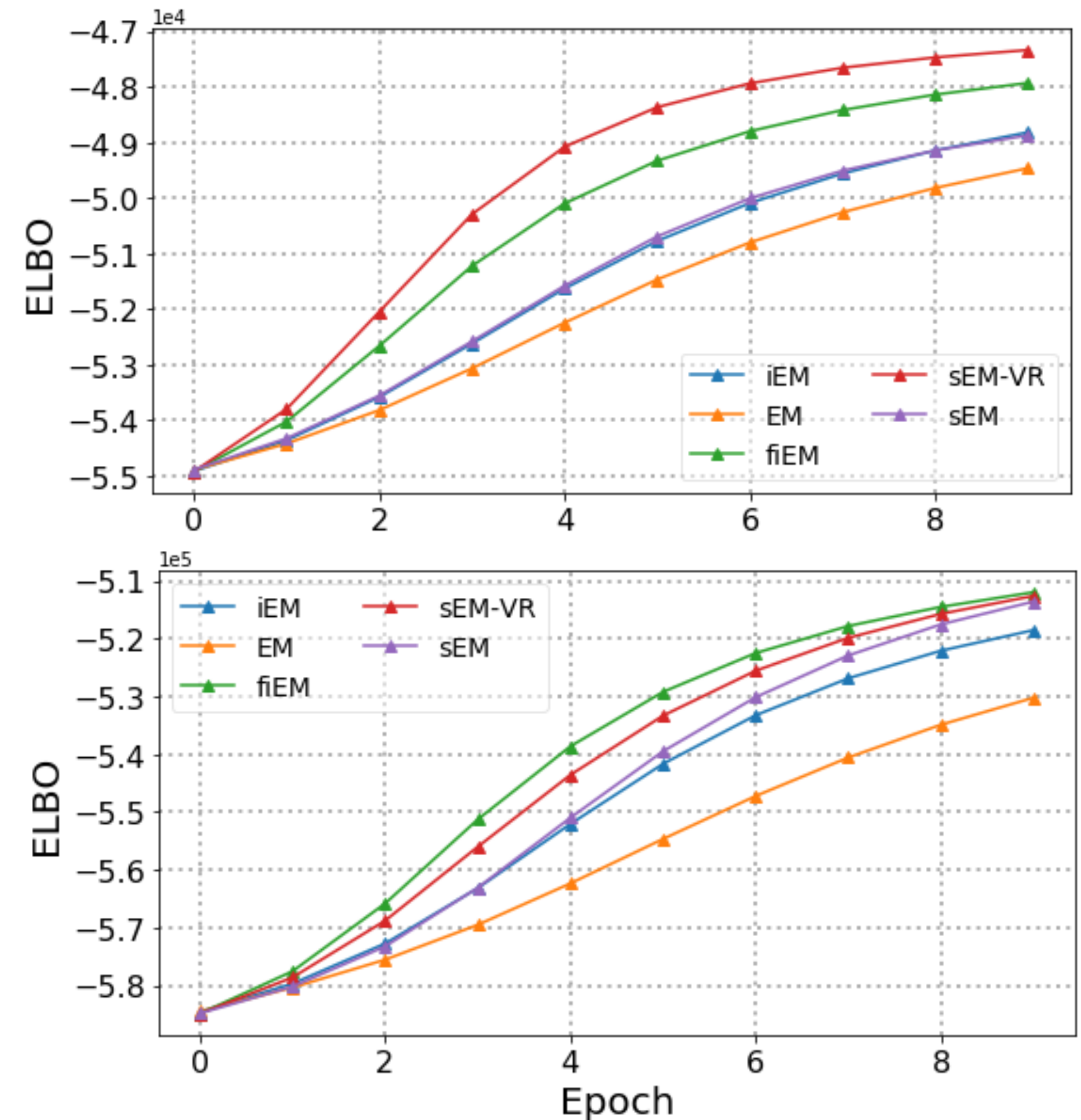
$$\begin{aligned} \log f(z_i, y_i; \boldsymbol{\theta}) &= \sum_{k=1}^K \sum_{d=1}^D \log(\boldsymbol{\theta}_{d,k}^{(t|d)}) \mathbb{1}_{\{k,d\}}(z_i, y_i^{(d)}) \\ &+ \sum_{k=1}^K \sum_{v=1}^V \log(\boldsymbol{\theta}_{k,v}^{(w|t)}) \mathbb{1}_{\{k,v\}}(z_i, y_i^{(w)}) \end{aligned}$$

- Penalization used:

$$R(\boldsymbol{\theta}^{(t|d)}, \boldsymbol{\theta}^{(w|t)}) = -\log \text{Dir}(\boldsymbol{\theta}^{(t|d)}; K, \alpha') - \log \text{Dir}(\boldsymbol{\theta}^{(w|t)}; V, \beta')$$

$$\boldsymbol{\theta} := (\boldsymbol{\theta}^{(t|d)}, \boldsymbol{\theta}^{(w|t)})$$

## Experiments





# Conclusion

# Take-Aways

- We studied the global convergence of stochastic EM Methods
  - Globally (independent of initialization)
  - Non-asymptotic results
- We used a Majorization-Minimization scheme to analyze the incremental EM method
- We interpreted the variance-reduced and the fast incremental method using a scaled gradient scheme to find a stationary point of a well defined Lyapunov function

**Thank You !**